

Korean Family Name Distribution in the Past

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We empirically study the genealogical trees of ten families for about five centuries in Korea. Although each family tree contains only the paternal part, the family names of women married to the family have been recorded, which allows us to estimate roughly the family name distributions for the past five hundred years. Revealed is the fact that the unique Korean family name distribution, characterized by a logarithmic form of the cumulative distribution and an exponentially decaying rank-size plot often called the Zipf plot, has remained unchanged for a long time. We discuss the implications of our results within a recently suggested theoretical model and compare them with observations in other countries in which power-law forms are abundantly found.

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I. INTRODUCTION

In the statistical-physics community, there has been intensive research interest in the form and the origin of various distributions observed in nature and in human societies [1]. To name a few, the degree distributions of many complex networks have been found to have power-law forms [2], and the word frequency in literature, the sizes of cities, the wealth distributions [3], the stock price returns [4], and magnitude of earthquakes have also been studied [1]. The family name distributions have also been a popular research arena for statistical physicists, and power-law forms have been observed in many countries [5,6] with some exceptions [7–9].

Systems of family names have been developed in most countries in order to make distinctions among families for various biological, sociological, and economical reasons: For instances, marriages among close family members have disastrous genetic effects, most people want to keep the wealth of the family within the family, and giving a social identity (you are one of us) to family members has huge advantages in farming, construction, and forming a bigger army. Once introduced, the inheritance of the family name in most countries follows a very simple rule: children's family names are from the father. However, the time evolution of the family name distribution is not yet completely understood. A recent study [8] has indicated the importance of the name generation rate in

explaining the very different family name distributions observed across countries. In particular, two groups of family name distributions have been found [8]: Korea [7] and China [9] exhibit exponentially decaying rank-size plots, or Zipf plots for the size versus the rank of families, and the number of family names increases logarithmically with population whereas for all other countries where the empirical studies have been made, both follow a power-law form.

In the present work, we use computerized data obtained from ten family books, each of which had been updated and inherited within a family from generation to generation. Each family book contained paternal genealogical trees; in addition, the names of women married to the family had been well recorded. In the analysis made here, we only use the names of married women with their birth years and their well-recorded family names (see Table 1). We then take snapshots of the family name distributions in the past, which reveals that the unique Korean name distribution has not changed for at least five hundred years.

This paper is organized as follows: In Sec. II, we report our main empirical results and discuss in Sec. III how different the observations are if we include information on the regional origins of family names. Comparisons with other countries and the implications of our results in a recently developed theoretical approach are discussed in Sec. IV.

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Table 1. Ten digitized family books, containing information on women married to the family, used in the present work [10]. Only women for whom their birth years and family names were properly recorded are used in analysis (N is the number of such women in each data set). Women do not change their family names after marriage in Korea. Here are also listed the numbers of family names (N_f) and those considering the regional origins (N_r).

data set	year begun	N	N_f	N_r
1	1513	104,366	165	2,668
2	1562	29,139	142	1,274
3	1439	17,911	121	923
4	1476	16,379	106	727
5	1698	15,445	125	915
6	1254	15,007	107	958
7	1475	11,526	112	736
8	1458	6,463	99	548
9	1752	3,500	89	390
10	1802	1,873	76	289

II. KOREAN FAMILY NAME DISTRIBUTION IN THE PAST

In Korean tradition, it is extremely unlikely that any Korean changes his (her) family name, which explains the existence of the small number of family names found in Korea (288 family names [7] in comparison to 132,000 names in Japan [6]). Such a small number of family names in Korea inevitably makes some families huge. For example, the biggest family name, Kim, covers more than 20 % of total population in Korea. Consequently, members in the huge-sized family cannot have a social group identity as a family, and classifications into sub-groups on finer scales have naturally been developed. More specifically, the regional origin of the family combined with the family name has been used to distinguish one family from the other (see Ref. 7 for more discussion); one of the authors of the present paper belongs to the sub-family group “Baek from Suwon,” with Suwon being the regional origin of his ancestor. We, in this Section, report our results for the case when the regional origins are disregarded; *i.e.*, families with the same name but with different regional origins are taken as one family. Observations made for other case with the regional origin taken into account are presented in the next Section.

In Figure 1, we display the rank-size plot for the number $n(r)$ of members, *i.e.*, the size of the family, versus the rank r of the family. For example, if a data point exists in the plot at $(r = 2, n = 10^3)$, it means that the second biggest family name has 10^3 members within it. It is remarkable to see in Figure 1(a) that all data sets, small and big, in Table 1 unanimously exhibit an exponential decay over a broad range of r . Although the total size of available data sets can cover only a tiny

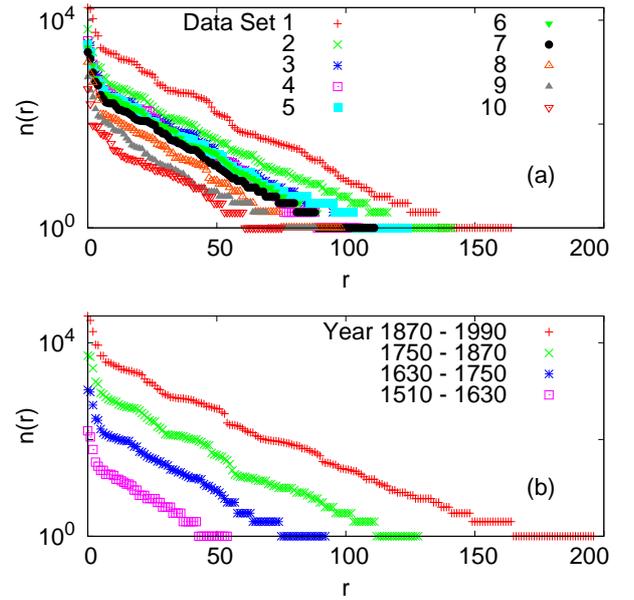


Fig. 1. (Color-online) (a) The size $n(r)$ versus the rank r of families (Zipf plot) taken from the 10 different data sets in Table 1. (b) All data sets in Table 1 are merged and then split according to birth years. It is clearly seen that (a) each data set exhibits qualitatively the same exponentially decaying rank-size plot and (b) the overall feature has not changed for about five hundred years.

fraction of the whole population within Korea for five hundred years, we strongly believe that the observation of the same exponentially decaying rank-size plot for all ten data sets can hardly be a coincidence. In Figure 1(b), we show the rank-size plot when all ten data sets are merged and then split into four different time periods, which confirms that the exponentially decaying rank-size plot has not changed for at least five hundred years in Korean history. Consequently, we suggest that the already found exponential rank-size plot in Ref. 7 from the census results for the years 1985 and 2000 has been a unique characteristic of the Korean family name distribution for a very long time since the family name system was introduced at least about two thousand years ago. Although limited by the available size of the family name information, it is also interesting to recognize that our method can provide a snapshot of the family name distribution in the past at any given time span.

Even if a rank-size plot looks similar to another as a whole, the actual order of frequent family names may be different for each data set. Suppose that we pick a data set and make a list of family names therein in a descending order of their family sizes. Then, the rank will naturally increase one by one along this list. In a different data set, however, the same list of family names will have a different order of ranks, in general. For example, we present Table 2 where the order becomes simply reversed. Denoting this new order as $\{r_i\}$ with $i = 1 \dots 5$, let us count P , the number of pairs (r_i, r_j) such that

Table 2. An example of the rank correlation between two data sets denoted as #1 and #2.

family name	A	B	C	D	E
ranks in #1	1	2	3	4	5
#2	5	4	3	2	1

Table 3. The Kendall tau rank correlation coefficients between the data sets used in the present work. Each element at the row i and the column j indicates $\tau(i, j)$, the Kendall tau rank correlation coefficient between the data sets i and j . Note that $\tau(i, j) = \tau(j, i)$, and that we consider only the family names commonly appearing in both of the data sets under comparison.

data set	1	2	3	4	5	6	7	8	9	10
1	1.00	0.78	0.73	0.84	0.77	0.77	0.64	0.82	0.79	0.77
2		1.00	0.75	0.76	0.74	0.70	0.61	0.70	0.66	0.70
3			1.00	0.71	0.75	0.73	0.67	0.71	0.73	0.74
4				1.00	0.76	0.75	0.58	0.80	0.74	0.77
5					1.00	0.70	0.62	0.80	0.73	0.72
6						1.00	0.63	0.75	0.78	0.81
7							1.00	0.61	0.65	0.65
8								1.00	0.78	0.77
9									1.00	0.75
10										1.00

$i < j$ and $r_i < r_j$. One can easily see the answer to be zero because the rank is only decreasing as we move to the right. On the other hand, if the new order had been identical with the original one, P would have amounted to $s(s - 1)/2$ with the number of elements of $s = 5$. Therefore the statistics P may be an indicator for the degree of rank correlation within the range between zero and $s(s - 1)/2$.

Normalizing P into $[-1, 1]$ gives the *Kendall tau rank correlation coefficient* as follows [11]:

$$\tau = \frac{2P}{s(s - 1)/2} - 1.$$

As discussed above, if the orders of ranks are unchanged across two data sets, τ equals unity while $\tau = -1$ for a perfectly reversed order. If rankings are independently given, τ will approach zero on average. Table 3 displays the rank correlations between the data sets in Table 1. The correlation between two different data sets is estimated as $\bar{\tau} = 0.73 \pm 0.06$, showing the significantly high correspondence among one another.

The functional form of the rank-size plot is, of course, closely related with the family name distribution function $P(k)$, which measures the frequency of the families of size k [7]. In particular, the exponential rank-size plot leads to the functional form $P(k) \sim k^{-1}$, which yields the logarithmic cumulative distribution function $P_{\text{cum}}(k) \sim -\ln k$. With the assumption of uniform sampling as discussed in Ref. 7, the number N_f of family names in a population of size N is expected to increase logarithmically, *i.e.*, $N_f \sim \ln N$. In comparison, for many other countries, the observed power-law behavior of $P(k) \sim k^{-\gamma}$ gives $P_{\text{cum}}(k) \sim k^{1-\gamma}$, and $N_f \sim N^{\gamma-1}$ is

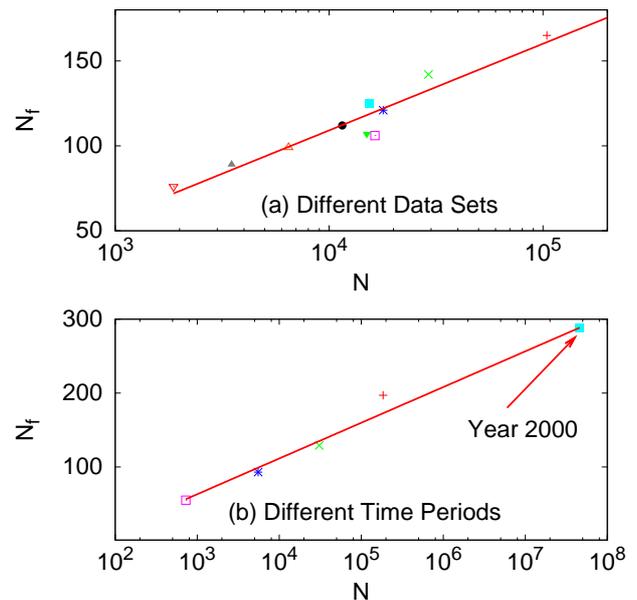


Fig. 2. (Color online) The number N_f of different family names without regional origins versus the population size N (a) for each date set in Table 1 and (b) for given time periods, in parallel to Figs. 1(a) and (b), respectively. Symbols in (a) and (b) correspond to the ones in Figure 1. It is shown clearly that $N_f \sim \ln N$ in (a). In (b), the actual result from the census in year 2000 is in accord with the logarithmic form.

expected. Consequently, the exponential rank-size plot in Figure 1 is expected to explain the logarithmic behavior in the number of family names, *i.e.*, $N_f \sim \ln N$, which is closely related with such a small number of fam-

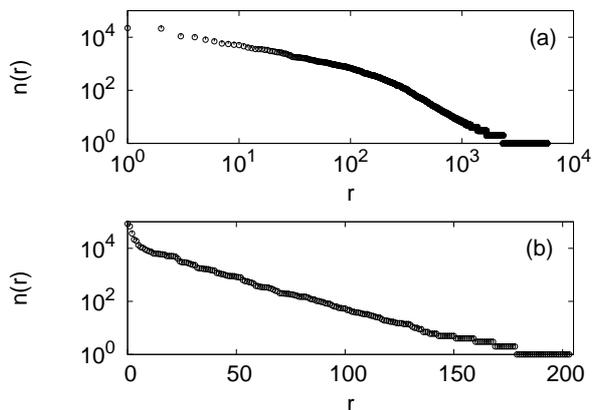


Fig. 3. Rank-size plots (a) with and (b) without regional origins. The rank-size plot without regional origins in (b) shows a very clear exponential decay as was observed in Figure 1, while it is much broader in (a) when the regional origins of family names are taken into account. All women’s names, regardless of the existence of the birth-year information, are included for better statistics.

ily names in Korea in comparison to other countries [8].

In order to validate the relation between the rank-size plot and $N_f(N)$, we display in Figure 2 the results corresponding to Figure 1. In Figure 2(a), N_f and N are taken from data sets in Table 1 with symbols corresponding to those in Figure 1(a), and the overall feature is well described by $N_f \sim \ln N$, as expected. In Figure 2(b), we again use the time-split data as in Figure 1(b), which is in accord with the logarithmic form, with the point for the year 2000 included. We again emphasize that our observation is in a sharp contrast to the finding in Japan [6], where N_f has been found to increase algebraically with N , not logarithmically.

III. FAMILY NAME DISTRIBUTION WITH REGIONAL ORIGINS

We next study the family name distribution in the past with the regional origins taken into account. In other words, we now distinguish “Kim from Gimhae,” to which one of the authors of the present paper belongs, from “Kim from Gyeongju,” and thus count the two as different family names. We believe that findings in Sec. II confirm the robustness of the family name distribution for a long time; accordingly, in order to make the sample size bigger, we include all women, regardless of the presence or absence of the birth-year information. In Figure 3, we summarize our results for the rank-size plot (a) with regional origins taken into account and (b) with regional origins disregarded, computed from all 420,719 women in comparison to 221,609 in Table 1 and in Sec. II. It is clearly seen that the rank-size plot again shows an

exponential decay form whereas the rank-size plot when the regional origins are included becomes very broad, presumably in a power-law form in the large r region. The total numbers of names with and without regional origins are 5,788 and 204, respectively.

IV. DISCUSSION

Recently, different family name distributions across countries have been explained as originating from different name generation behaviors [8]. In detail, from the use of the master equation approach, we have shown that in a society where new family names are rarely invented, the rank-size plot is expected to be of an exponential form. On the other hand, if the new name generation rate is proportional to the population size, the rank-size plot is of the power-law form. We in this work have shown that in the past five hundred years in Korea, the exponential rank-size plot has been unchanged, which is closely related with such a small number of family names in Korea according to Ref. 7.

In summary, we have used computerized family genealogical tree information for ten families to extract the birth years and the names of women who were married to each family. Clearly shown is the robustness of the family name distribution in Korea for the past five hundred years, at least. Particularly, the rank-size plot has been shown to be of an exponential form, which confirms the finding in Ref. 7 made from the results of the recent census. This is in itself an interesting surprise because we expected that in Korean history, families in a higher social class must have chosen women in equally influential families, distorting the women’s family name distributions. Additionally, our empirical explorations have also indicated that the family name distributions with and without regional origin information of married women are very different: The latter takes the exponential form, and the former has a much broader decay form, in close resemblance to the family name distributions in other countries.

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REFERENCES

- [1] See, M. E. J. Newman, *Contemp. Phys.* **46**, 323 (2005) and references therein.
- [2] R. Albert and A.-L. Barabási, *Rev. Mod. Phys.* **74**, 47 (2002); S. N. Dorogovtsev and J. F. F. Mendes, *Adv. Phys.* **51**, 1079 (2002); M. E. J. Newman, *SIAM Rev.* **45**, 167 (2003); S.-W. Son and H. Jeong, *J. Korean Phys. Soc.* **48**, S208 (2006).
- [3] G. Lee and G. I. Kim, *J. Korean Phys. Soc.* **49**, 1675 (2006).
- [4] G. Oh and S. Kim, *J. Korean Phys. Soc.* **48**, S197 (2006).
- [5] J. E. Dipierri, E. L. Alfaro, C. Scapoli, E. Mamolini, A. Rodriguez-Larralde and I. Barrai, *Am. J. Phys. Anthropol.* **128**, 199 (2005); C. Scapoli, E. Mamolini, A. Carri-eri, A. Rodriguez-Larralde and I. Barrai, *Theor. Popul. Biol.* **71**, 37 (2007); D. H. Zanette and S. C. Marunbia, *Physica A* **295**, 1 (2001); W. J. Reed and B. D. Hughes, *ibid.* **319**, 579 (2003).
- [6] S. Miyazima, Y. Lee, T. Nagamine and H. Miyajima, *Physica A* **278**, 282 (2000).
- [7] B. J. Kim and S. M. Park, *Physica A* **347**, 683 (2005).
- [8] S. K. Baek, H. A. T. Kiet and B. J. Kim, *Phys. Rev. E*, in press.
- [9] Y. Yuan *Chinese Surnames* (East China Normal University Press, Sanghai, 2002), p. 21.
- [10] In order to protect the privacy of individuals, only the family name field from the family book has been provided for data analysis.
- [11] H. Abdi, *The Kendall Rank Correlation Coefficient*, edited by N. J. Salkind *Encyclopedia of Measurement and Statistics* (Sage, Thousand Oaks, CA, 2007).